Biobehavioral, Immune, and Health Benefits following Recurrence for Psychological Intervention Participants

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Abstract

Purpose: A clinical trial was designed to test the hypothesis that a psychological intervention could reduce the risk of cancer recurrence. Newly diagnosed regional breast cancer patients (n = 227) were randomized to the intervention-with-assessment or the assessment-only arm. The intervention had positive psychological, social, immune, and health benefits, and after a median of 11 years the intervention arm was found to have reduced the risk of recurrence (hazard ratio, 0.55; P = 0.034). In follow-up, we hypothesized that the intervention arm might also show longer survival after recurrence. If observed, we then would examine potential biobehavioral mechanisms.

Experimental Design: All patients were followed; 62 recurred. Survival analyses included all 62. Upon recurrence diagnosis, those available for further biobehavioral study were accrued (n = 41, 23 intervention and 18 assessment). For those 41, psychological, social, adherence, health, and immune (natural killer cell cytotoxicity, T-cell proliferation) data were collected at recurrence diagnosis and 4, 8, and 12 months later.

Results: Intent-to-treat analysis revealed reduced risk of death following recurrence for the intervention arm (hazard ratio, 0.41; P = 0.014). Mixed-effects follow-up analyses with biobehavioral data showed that all patients responded with significant psychological distress at recurrence diagnosis, but thereafter only the intervention arm improved (P values < 0.023). Immune indices were significantly higher for the intervention arm at 12 months (P values < 0.017).

Conclusions: Hazards analyses augment previous findings in showing improved survival for the intervention arm after recurrence. Follow-up analyses showing biobehavioral advantages for the intervention arm contribute to our understanding of how improved survival was achieved. Clin Cancer Res; 16(12):3270–8. ©2010 AACR.

Meta-analyses suggest that stress-related psychosocial factors (1) and lower health-related quality of life (2) are associated with poorer cancer-related survival, with a 13% increase in the hazard ratio (HR) in studies of breast cancer patients (1). In 1994 a randomized controlled trial, the Stress and Immunity Breast Cancer Project (SIBCP), was designed to test the hypothesis that newly diagnosed breast cancer patients receiving a psychological intervention would have a reduced risk of recurrence and breast cancer death compared with patients who were only assessed. A conceptual model guided the development of the clinical trial. The Biobehavioral Model of Cancer Stress and Disease Course (3) proposes that psychological stress leads to disruptions in quality of life, health behaviors, and immunity, which in turn contribute to poorer disease outcomes. It was hypothesized that an intervention designed to reduce emotional distress and improve social adjustment, health behaviors, and adherence might also improve immunity and disease course. Analyses showed that positive intervention effects were achieved across the psychological and immune outcomes at 4 months (4) with similar effects and improved health at 12 months (5). Moreover, recent end point analyses show that after a median follow-up of 11 years, the intervention arm had a reduced risk of breast cancer recurrence (HR, 0.55; P = 0.034) and breast cancer death (HR, 0.44; P = 0.016) compared with the assessment-only arm (6).

As cancer progresses, the ability of psychological intervention to affect disease course may change. The intervention was hypothesized to reduce the risk of recurrence, and it may have been solely through this reduction in recurrence risk that mortality was affected.
**Translational Relevance**

This article describes late translational research examining the long-term benefits of a psychological intervention for newly diagnosed breast cancer patients. Based on a theoretical model, the Biobehavioral Model of Cancer Stress and Disease Course, the intervention sought to reduce stress and distress, improve social support, improve health behaviors (including medical adherence), and ultimately improve disease course. The intervention was delivered using a session-by-session manual with the intent that the intervention could be easily disseminated. Thus, this psychological intervention could be incorporated into a comprehensive approach to cancer care.

However, if the intervention continued to affect psychological and immune functioning after the recurrence, it is plausible that survival benefits would have also persisted. Therefore, we tested whether intervention patients also showed a reduced risk of breast cancer death after recurrence. Of course, recurrence is devastating, with our previous studies showing patients’ stress to be equivalent to that reported at the time of the initial diagnosis (7, 8). Patients who do not have the benefit of a psychological intervention at recurrence may have depressive symptoms and anxieties (9, 10), and we anticipated that the assessment-only patients would respond similarly. In fact, we anticipated that all patients would be significantly distressed when learning of their recurrence. If, however, there were residual benefits from having earlier learned strategies for reducing stress and enhancing coping, the intervention arm patients might evidence improved adaptation.

In summary, we first tested for study arm differences in survival following recurrence. We then tested for study arm differences in biobehavioral variables to determine if the intervention arm evidenced more favorable trajectories than the assessment-only arm did in the 12 months following recurrence.

**Patients and Methods**

**Description of the SIBCP trial**

Complete descriptions of accrual, sample characteristics, stratification and randomization, power estimates, assessment, intervention, and follow-up procedures have been published (4). Briefly, patients \( n = 227 \) newly diagnosed with regional breast cancer and awaiting the start of adjuvant therapy after surgery were accrued. The majority had stage II (90%) rather than stage III (10%) disease, were estrogen/progesterone receptor (ER/PR) positive (68%), and were premenopausal (54%). The majority (83%) received their treatment at the Ohio State University National Cancer Institute–designated Comprehensive Cancer Center (OSUCCC). As previously reported (4), the intervention–with assessment \( (n = 114) \) and the assessment-only \( (n = 113) \) arms did not differ with regard to sociodemographic characteristics, extent of disease, prognostic factors, or cancer treatments received \( (P \text{ values} > 0.23) \).

The intervention was conducted in groups of 8 to 12 patients with two clinical psychologist leaders. It included relaxation training, positive ways to cope with stress and cancer-related difficulties (e.g., fatigue), methods to maximize social support, and strategies for improving health behaviors (diet, exercise) and adherence to cancer treatments. A total of 26 sessions (39 therapy hours) were delivered over 12 months (4). A detailed description is available (11). Patients were reassessed every four months during year one, every six months during years 2 to 5, and annually thereafter.

**Patients and procedures for recurrence substudy**

As patients recurred, they were approached for accrual (Fig. 1). Recurrence was defined as the development of breast cancer in the treated breast or chest wall, or at a distant site. Second primary tumors (e.g., contralateral breast, endometrial) did not meet this criterion. Informed consent was obtained, with study approval from the local institutional review boards and the Department of Health and Human Services assurances.

By October of 2007, 62 patients had recurred. Of them, 17 (27%) were not available for substudy accrual; 4 patients (6%) progressed rapidly and died, and 13 (21%) had previously discontinued SIBCP assessments. Of the remaining 45, 4 (6%) patients declined participation. For the 41 accrued, biobehavioral assessments were conducted a median of 11 weeks after diagnosis (baseline) and 4, 8, and 12 months later. Although biobehavioral data are available only for 41, sociodemographic, recurrence disease and treatments, and survival data are available for all 62 who recurred.

**Measures**

**Psychological distress.** The Profile of Mood States-Short Form (POMS; refs. 12, 13) is a 37-item inventory assessing negative mood. A total mood disturbance score which ranges from -24 to 124 was used. Internal consistency was 0.92.

**Social support.** The Interpersonal Support Evaluation List (ISEL; ref. 14) assessed patients' general perception of support. Scores range from 6 to 24. Internal consistency was 0.90. The Perceived Social Support Scales (PSS; ref. 15) assessed support from friends (PSS-Fr) and family (PSS-Fa) and was administered at baseline only. Scores range from 0 to 20 for each. Internal consistency was 0.85 and 0.92, respectively.

**Adherence: chemotherapy dose intensity.** Dose intensity (DI) for chemotherapy regimens was calculated, as recommended (16). Six assessment and 13 intervention patients received chemotherapy in the 12 months following recurrence diagnosis, with common agents being capecitabine \( (n = 5) \), cisplatin/carboplatin \( (n = 5) \), and the taxanes \( (n = 13) \).
**Immune assays.** Procedures for blood separation, quantification of T lymphocytes and natural killer (NK) and T-cell subsets, and immune assay procedures have been detailed (4, 5). Natural killer cell cytotoxicity (NKCC) was expressed as the mean of standardized scores (Z-scores) from six E:T ratios, as previously described (17). Blastogenesis responses were expressed as the mean of standardized scores from each of the three dilutions of concanavalin A (Con A) and phytohemagglutinin (PHA) as previously described (17).

**Health.** A research nurse assessed patients. Items for evaluating symptoms, signs, illnesses, lab values, etc. (symptoms/signs), came from the Southwest Oncology Collaborative Group (1994 version) toxicity measure, as described (5). Karnofsky performance status (KPS; ref. 18) was rated for use as a control variable in the analyses.

**Other.** Patients in both arms were queried about their receiving any psychotherapy or counseling as previously reported (6). The same questions were used with the addition of 15 others assessing use of complementary/alternative therapies (e.g., yoga, energy healing).

**Statistical analysis**

**Survival analysis.** Descriptive data are provided (n = 62). The prior SIBCP hazards analyses (n = 227) showed advantages for the intervention arm in risk of recurrence, defined as time from randomization to biopsy/study confirming first recurrence (6). The current analyses tested the period from biopsy/study confirming first recurrence to breast cancer death. Cox proportional hazards analysis (19) was used to test for study arm differences. Known prognostic factors (20) were examined as covariates. From the initial cancer diagnosis, we used cancer stage, ER/PR status, tumor grade, adjuvant therapy received, and measures with a significant group difference at study baseline (KPS, POMS; ref. 4); and from the recurrence diagnosis, we used disease-free interval (time from randomization to biopsy/study confirming first recurrence), site of recurrence (locoregional versus distant), and type of cancer treatment received after recurrence. Of these possible covariates, if there were variables highly correlated with each other (r > 0.6), only one was considered as a covariate to avoid multicollinearity. For model development, a backward elimination procedure was used as it has been recommended as one of the best methods for identifying important covariates when there are a large number of covariates to be considered with a relatively small number of events (21, 22). The covariates that predicted the end point with P < 0.1 were retained in the final model. The

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Fig. 1. Study flow from the time of randomized controlled trial accrual through a median of 11 years of follow-up. Patients designated as “lost follow-up” are known to be alive; cancer status, however, is not known.
log linearity assumption for continuous covariates was tested using the procedures recommended by Hosmer, Lemeshow, and May (23). The proportional hazards assumption was tested using the log minus log test.

**Biobehavioral analysis.** Preliminary analyses were conducted for the biobehavioral analyses. First, analyses contrasted the biobehavioral study participants (n = 41) and nonparticipants (n = 21). Second, ANOVA or $\chi^2$, as appropriate, compared participants in the assessment-only (n = 18) and the intervention (n = 23) arms on sociodemographic, prognostic, and treatment variables. For the biobehavioral variables, mixed-effects modeling (24) tested for study arm differences in baseline level and the trajectory of change following recurrence. Control variables were considered. To reduce variability due to cancer treatments, surgery, chemotherapy, and radiation and the treatment by time interactions were controlled in all models when relevant for specific assessments (e.g., surgery at baseline, chemotherapy at four months). In addition, baseline NK cell counts for the NKCC model and baseline total T-cell counts for the Con A and PHA models were included to control for absolute cell numbers in the analysis of functional responses. Nonsignificant control terms ($P > 0.1$) were removed from each model until a final solution was reached, as recommended (25).

For these analyses, the study arm and the study arm $\times$ time effects were of primary interest. The study arm effect indicates a group difference in intercept, testing whether study arms differed at baseline (recurrence diagnosis). The study arm $\times$ time effect tests whether the rate of change differed between arms, with the prediction being more positive adaptation and outcomes for the intervention arm. Bonferroni’s adjustment for multiple comparisons was used for the immune outcomes (i.e., reduction in $\alpha$ level to 0.017). Effect sizes, partial correlation coefficients ($pr$) computed using $t$ values and degrees of freedom (26), are provided. All statistical tests were two-sided.

**Table 1.** Equivalence of the assessment-only and intervention arms on sociodemographic, prognostic, and cancer treatments received variables ($n = 62$)

<table>
<thead>
<tr>
<th>Variables</th>
<th>Assessment only ($n = 33$)</th>
<th>Intervention ($n = 29$)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mean (SD)/%</td>
<td>Mean (SD)/%</td>
</tr>
<tr>
<td><strong>Sociodemographic variables at recurrence</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age (y)</td>
<td>52.2 (10.8)</td>
<td>55.5 (11.9)</td>
</tr>
<tr>
<td>Race (% Caucasian)</td>
<td>94</td>
<td>90</td>
</tr>
<tr>
<td>Education (y)</td>
<td>14.0 (2.3)</td>
<td>14.9 (2.7)</td>
</tr>
<tr>
<td>Marital status (% married)</td>
<td>67</td>
<td>66</td>
</tr>
<tr>
<td>Family income (thousand $/year)</td>
<td>51.6 (39.7)</td>
<td>54.2 (40.6)</td>
</tr>
<tr>
<td><strong>Initial diagnosis</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Stage (I vs. III, % II)</td>
<td>91</td>
<td>86</td>
</tr>
<tr>
<td>Nodes (number positive)</td>
<td>4.8 (7.3)</td>
<td>4.9 (9.1)</td>
</tr>
<tr>
<td>ER/PR (% positive)</td>
<td>55</td>
<td>55</td>
</tr>
<tr>
<td><strong>Treatments for initial diagnosis</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Surgery (% modified radical mastectomy)</td>
<td>61</td>
<td>62</td>
</tr>
<tr>
<td>Chemotherapy (% yes)</td>
<td>85</td>
<td>83</td>
</tr>
<tr>
<td>Radiation therapy (% yes)</td>
<td>49</td>
<td>45</td>
</tr>
<tr>
<td>Hormonal therapy (% yes)</td>
<td>73</td>
<td>62</td>
</tr>
<tr>
<td><strong>Recurrence diagnosis</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Median disease-free interval (mo)</td>
<td>26.0</td>
<td>33.0</td>
</tr>
<tr>
<td>Site of recurrent disease (%)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Local</td>
<td>29</td>
<td>18</td>
</tr>
<tr>
<td>Regional</td>
<td>0</td>
<td>7</td>
</tr>
<tr>
<td>Distant</td>
<td>71</td>
<td>75</td>
</tr>
<tr>
<td>Bone only</td>
<td>45</td>
<td>33</td>
</tr>
<tr>
<td>Distant, other</td>
<td>55</td>
<td>67</td>
</tr>
<tr>
<td><strong>Treatments for recurrence diagnosis</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Surgery (% yes)*</td>
<td>43</td>
<td>14</td>
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<tr>
<td>Chemotherapy (% yes)</td>
<td>47</td>
<td>57</td>
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<tr>
<td>Radiation therapy (% yes)</td>
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<tr>
<td>Hormonal therapy (% yes)</td>
<td>47</td>
<td>50</td>
</tr>
<tr>
<td>Bone marrow transplantation (% yes)</td>
<td>3</td>
<td>11</td>
</tr>
</tbody>
</table>

*Significant group difference, $P < 0.05$. 

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**Results**

**Survival analyses**

As of October 2007, 62 recurrence cases (29 intervention arm and 33 assessment arm) were confirmed through patient contact or cause of death determination from death certificates. Table 1 provides descriptive information. Patients (n = 62) were primarily middle aged (mean, 54 years; SD, 11), with some college (69%), and married/partnered (66%). Caucasians accounted for 92%, African Americans for 8%. At initial diagnosis, the majority (89%) had stage II disease. At recurrence diagnosis, the median disease-free interval was 30 months (range, 2-144 months) and the majority (73%) had distant rather than locoregional metastases. The majority (90%) were treated at OSUCCC for their recurrence, receiving surgery (29%), chemotherapy (52%), radiation therapy (29%), hormonal therapy (48%), and/or bone marrow transplantation (7%).

Of the 62 who recurred, 44 (71%; 19 intervention arm and 25 assessment arm) patients died, all from breast cancer. Median time to death was 2.8 years (range, 0.9-11.8 years) for the intervention arm and 2.2 years (range, 0.2-12.0 years) for the assessment arm. Multivariate comparison with the Cox proportional hazards analysis showed a significant effect of study arm adjusting for initial functional status, psychological distress, and recurrence site: patients in the intervention arm had a lower risk of death after recurrence (HR, 0.41; 95% confidence interval, 0.20-0.83; P = 0.014; Table 2 and Fig. 2).

**Biobehavioral analyses**

**Preliminary analyses.** Two comparisons were made. First, analyses contrasted patients who recurred and participated in the follow-up (n = 41) with patients who recurred but did not participate (n = 21). The groups were not significantly different in sociodemographic (education, marital status, income, employment status), prognostic (stage, number of nodes positive, ER/PR status, menopausal status), previous cancer treatments received (surgery type, chemotherapy type and dose, radiation therapy, hormonal therapy), or disease-free interval variables (P values > 0.128). The only difference was age (P = 0.017), with participants being older when diagnosed initially (mean age, 53 years) than the nonparticipants (mean age, 46 years). The current sample might underestimate the patients' psychological distress as previous research (27) indicated a faster quality of life recovery after recurrence among older patients (age ≥54 years) compared with younger patients.

For the participants (n = 41), analyses contrasted the assessment (n = 18) and the intervention (n = 23) arms. There were no significant differences in sociodemographic, prognostic, and previous/current cancer treatments received variables (P values > 0.059). The only difference between arms was receiving of surgery at recurrence (P = 0.004). The higher rate of surgery in the assessment arm reflects the higher rate of local recurrence, although differences between groups in site of recurrence was not statistically significant. For patients receiving surgery, surgical sites were as follows: for intervention, all three received breast surgery; for assessment, seven received breast, two brain, and one lung surgery. Therefore, receiving of surgery was controlled in the analyses regardless of significance. Lastly, study arms did not differ in their

**Table 2. Final multivariate Cox proportional hazards model contrasting study arms in survival after recurrence (n = 62)**

<table>
<thead>
<tr>
<th>Variable</th>
<th>HR (95% CI)</th>
<th>P</th>
</tr>
</thead>
<tbody>
<tr>
<td>Study arm (intervention vs. assessment only)</td>
<td>0.410 (0.202-0.832)</td>
<td>0.014</td>
</tr>
<tr>
<td>Initial functional status (KPS)</td>
<td>0.346 (0.170-0.704)</td>
<td>0.003</td>
</tr>
<tr>
<td>Initial psychological distress (POMS)</td>
<td>0.982 (0.967-0.997)</td>
<td>0.022</td>
</tr>
<tr>
<td>Recurrence site (locoregional vs. distant)</td>
<td>0.191 (0.080-0.458)</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

Abbreviation: 95% CI, 95% confidence interval.
involvement in any type of counseling at any time (P values > 0.457). Of the 15 complementary therapies surveyed, study arms did not differ in use of the most commonly used (i.e., spiritual healing, megavitamin therapy, dietary changes) or any others (P values > 0.147).

**Main analyses.** Descriptive statistics for all secondary measures at baseline and 12 months are available as supplementary information. Table 3 presents all mixed-effects model outcomes comparing study arms in biobehavioral trajectories during the 12 months following recurrence diagnosis.

For psychological distress, there was no significant study arm effect (P = 0.430). As predicted, both arms responded to the diagnosis of recurrence with high levels of negative mood (POMS). The time effect was not significant (P = 0.523), indicating no reduction in negative mood from baseline to 12 months for the assessment arm. However, the intervention arm showed a decline in negative mood (i.e., emotional improvement), and the study arm × time effect indicated that the difference between the groups was significant (P = 0.008; Fig. 3A).

For social support, there was a significant study arm effect (P = 0.001). The time effect was significant (P = 0.002), with the assessment arm showing a decline (i.e., less social support). The nonsignificant study arm × time effect (P = 0.456) indicates that the baseline group difference – significantly higher levels of social support for the intervention arm – was maintained (Fig. 3B). Results for the PSS scales using analysis of covariance were consistent. Intervention arm patients perceived significantly more support at recurrence diagnosis from their family (PSS-Fa; P = 0.007). The effect for PSS-Fr was similar, though P = 0.086.

Chemotherapy DI was compared using ANOVA. The groups differed significantly (P = 0.024), with a mean regimen DI of 94% (SD, 9; range, 74-100) for the intervention arm and 81% (SD, 11; range, 63-92) for the assessment arm. The significant group difference and the large effect size (Cohen’s d, 1.2) are notable, but they require replication to determine their reliability. Moreover, the relevance of DI to survival following breast cancer recurrence is unclear (28, 29).

For health, there was no significant study arm or study arm × time effect (P values > 0.614). Both arms showed a significant reduction in symptoms/signs with time (P values < 0.034).

With regard to immunity, there was no significant study arm effect (P = 0.771) for NKCC. The time effect indicated

<table>
<thead>
<tr>
<th>Variable</th>
<th>Effect</th>
<th>Estimate (95% CI)</th>
<th>t</th>
<th>Pr</th>
</tr>
</thead>
<tbody>
<tr>
<td>Psychological distress</td>
<td>Study arm</td>
<td>5.811 (−8.896 to 20.518)</td>
<td>0.798</td>
<td>0.12</td>
</tr>
<tr>
<td></td>
<td>Time</td>
<td>0.253 (−0.553 to 1.058)</td>
<td>0.649</td>
<td>0.05</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>−1.282 (−2.195 to -0.370)</td>
<td>−2.908*</td>
<td>−0.26</td>
</tr>
<tr>
<td>Social support</td>
<td>Study arm</td>
<td>2.800 (1.336-4.264)</td>
<td>3.925*</td>
<td>0.60</td>
</tr>
<tr>
<td></td>
<td>Time</td>
<td>−0.131 (−0.206 to −0.056)</td>
<td>−3.710*</td>
<td>−0.68</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>0.027 (−0.049 to 0.102)</td>
<td>0.769</td>
<td>0.22</td>
</tr>
<tr>
<td>Health</td>
<td>Study arm</td>
<td>0.019 (−0.056 to 0.094)</td>
<td>0.508</td>
<td>0.08</td>
</tr>
<tr>
<td></td>
<td>Time</td>
<td>−0.004 (−0.008 to −0.0004)</td>
<td>−2.208*</td>
<td>−0.35</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>−0.001 (−0.005 to 0.002)</td>
<td>−0.694</td>
<td>−0.14</td>
</tr>
<tr>
<td>Immunity</td>
<td>Study arm</td>
<td>0.071 (−0.427 to 0.568)</td>
<td>0.294</td>
<td>0.06</td>
</tr>
<tr>
<td></td>
<td>Time</td>
<td>−0.169 (−0.249 to −0.089)</td>
<td>−4.292†</td>
<td>−0.60</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>0.101 (0.024-0.178)</td>
<td>2.701†</td>
<td>0.47</td>
</tr>
<tr>
<td>'A</td>
<td>Study arm</td>
<td>−0.030 (−0.631 to 0.570)</td>
<td>−0.104</td>
<td>0.02</td>
</tr>
<tr>
<td>Con A</td>
<td>Time</td>
<td>−0.064 (−0.146 to 0.018)</td>
<td>−1.587</td>
<td>−0.27</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>0.068 (−0.030 to 0.165)</td>
<td>1.411</td>
<td>0.24</td>
</tr>
<tr>
<td>PHA</td>
<td>Study arm</td>
<td>0.476 (−0.081 to 1.033)</td>
<td>1.766</td>
<td>0.35</td>
</tr>
<tr>
<td></td>
<td>Time</td>
<td>−0.016 (−0.092 to 0.059)</td>
<td>−0.456</td>
<td>−0.11</td>
</tr>
<tr>
<td></td>
<td>Study arm × time</td>
<td>0.049 (−0.038 to 0.136)</td>
<td>1.192</td>
<td>0.28</td>
</tr>
</tbody>
</table>

*P < 0.05.
†P < 0.017 with Bonferroni’s adjustment.
A significant decline in NKCC for the assessment arm \((P < 0.001)\). The study arm × time effect was significant \((P = 0.012)\), with the intervention arm showing a slower rate of change than the assessment arm (Fig. 3C). For Con A and PHA blastogenesis, none of the effects was significant \((P \text{ values} > 0.090)\). As immunity was included in the trial as a potential mechanism for improved disease course (3), a follow-up analysis determined if the immune trajectories led to differential levels at 12 months. The same mixed-effects model contrasted the study arms with the intercept coded to be the 12-month assessment. Results showed that the intervention arm had significantly higher NKCC \((P = 0.001)\), Con A blastogenesis \((P = 0.021)\), and PHA blastogenesis \((P = 0.007)\) at 12 months.

**Discussion**

Hypotheses were tested to illuminate survival effects from the SBCP trial. The prior Cox proportional hazards analysis had shown that the intervention arm had a 45% reduced risk for breast cancer recurrence (6). The current analysis shows that following recurrence, the intervention arm had a 59% reduction in the risk of dying of breast cancer. In combination, the data suggest that factors or mechanisms impacting disease progression occurred during both intervals: mechanisms influential from initial diagnosis to recurrence and mechanisms influential from recurrence to breast cancer death.

Collection of biobehavioral data following recurrence was planned shortly after the trial began. A majority of the recurrence patients (66%) were studied, and the availability of repeated assessments increased power and reliability for the analyses. The mixed-effects model analysis uses all available data, which enhances generalizability. Even considering these factors, we cannot be certain, of course, that identical biobehavioral data patterns would be observed with a sample size of 62 when compared with the 41. However, all findings were in the predicted direction, replicating, and in some cases extending, the effects observed at the time of intervention delivery (4, 5) and with larger effect sizes.

For breast cancer patients, recurrence may be particularly stressful because currently available treatments are not curative (30), posing a significant therapeutic challenge for the oncologist (31). As in previous studies (7, 27), all patients reported negative moods immediately following diagnosis. Thereafter, study arms diverged as they had when the intervention was delivered (4, 5), with emotional distress significantly declining for the intervention arm, yet remaining high for the assessment arm. A plausible hypothesis for this pattern is that after the shock of diagnosis, the intervention arm patients may have drawn upon strategies they had learned for reducing stress and coping positively. For them, a median of 18 months had passed from the conclusion of intervention to recurrence diagnosis. Importantly, ancillary data show that the decline in distress for the intervention arm was not due to differential receiving of other counseling or complementary therapies.

Intervention arm patients also seem to have been more successful at securing and maintaining social support. Significant gains in perceived social support had been found for the intervention arm during intervention delivery and...
(4). This may have led to stable changes in their most important relationships, or at a minimum, changes helpful during times of stress. Upon recurrence, the social support advantage for the intervention patients never diminished, and the lower support level reported by the assessment patients never improved (and actually declined).

Regarding the immunity findings, our previous studies had shown that as patients entered the trial, stress covaried with lower levels of NKCC and T-cell blastogenesis (32). During the intervention, T-cell blastogenesis was significantly higher for the intervention arm (4, 5). It remains to be seen whether immunity as measured here is critical to survival of breast cancer patients. High levels of NKCC covary with a better prognosis and overall survival in many different cancers (33–35). Impaired NKCC has been found to correlate with stage of disease (36), and NK cell infiltration into primary tumors is associated with less lymphatic invasion, fewer metastases, and better survival (37).

The relationship between the psychological intervention effects and breast cancer survival will require further study. In the general case, however, clinical studies suggest that tumor cells remain dormant for varied periods of time following surgery and adjuvant chemotherapy, some of which are long (38). The prevailing theory suggests that tumor cells may be quiescent and/or at equilibrium with the patient’s immune system. In support of this concept are studies showing that the maintenance of immune surveillance is important for long-term suppression of cancer cell growth. In these models, the immune system is able to control tumor growth, although it is unable to eliminate all of the tumor cells (39, 40). Theoretically, a cancer recurrence might occur and become clinically detectable when tumor cells overcome the multiple immune mechanisms (CD8+ T cells, NK cell surveillance, antibodies) that can inhibit tumor growth. Once tumor cells escape immune surveillance, immune effector cell function may be further impaired due to tumor-derived factors. For example, NK cell maturation in the bone marrow and NK cell recruitment to the tumor site are inhibited as tumor burden increases (41, 42). The latter is consistent with Pierson and Miller’s (43) comparison of cancer patient and healthy donor samples, which suggested both NK cell number and NKCC are distinctly and negatively affected by the malignant environment.

In summary, intent-to-treat analyses show improved survival for SIBC intervention arm patients following recurrence. Other recent trials have reported null survival effects when providing an intervention to patients recently diagnosed with metastatic breast cancer (44–46). These trials and a similar one (47) all used a therapy (supportive expressive; ref. 48) that differs from that of the SIBC (11). Reviews of the literature (49, 50) show that psychological interventions are typically delivered at the time of initial diagnosis, and it has not been considered that patients might derive benefits from them much later or at the time of recurrence. The findings suggest that if psychological interventions are offered early, they may provide enduring, late benefits and possibly longer survival.

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No potential conflicts of interest were disclosed.

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Correction: Biobehavioral, Immune, and Health Benefits following Recurrence for Psychological Intervention Participants

In this article (Clin Cancer Res 2010;16:3270–8), which was published in the June 15, 2010, issue of *Clinical Cancer Research* (1), there was an error in the second paragraph of the Results section. It currently reads, "Median time to death was 2.8 years (range, 0.9–11.8 years) for the intervention arm and 2.2 years (range, 0.2–12.0 years) for the assessment arm." These data were correctly reported previously as the median time to recurrence (2). The sentence should read, "Median time to death was 3.2 years (range, 0–6.6 years) for the intervention arm and 1.7 years (range, 0.2–4.2 years) for the assessment arm." These data suggest a doubling of time to the endpoint. The authors regret any confusion.

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